

This is the peer reviewed version of the following article: Griffin, R. M., Hayward, A. D., Bolund, E. , Maklakov, A. A., Lummaa, V. and Gaillard, J. (2018), Sex differences in adult mortality rate mediated by early-life environmental conditions. *Ecology Lett*, 21: 235-242, which has been published in final form at <https://doi.org/10.1111/ele.12888>. This article may be used for non-commercial purposes in accordance With Wiley Terms and Conditions for self-archiving.

1 **Title**

2 Sex differences in adult mortality rate mediated by early-life environmental conditions

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16 **Running Title**

17 Ecology of sex differences in adult mortality

18 **Keywords**

19 Development, Environmental Variation, Humans, Life-History, Sexual Dimorphism

20 **Article Type**

21 Letter

22 **Counts**

23 Abstract = 150 words

24 Main text = 3900 words

25 References = 34

26 Figures = 4 (+4 supplementary)

27 Tables = 0 (+2 supplementary)

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30 **Statement of Authorship**

31 RMG and VL designed the study, VL collected data, RMG analysed data, all authors contributed  
32 to writing of the manuscript.

33 **Data Accessibility Statement**

34 Data will be made available via Dryad upon acceptance

35

36 **Abstract**

37 Variation in sex differences is affected by both genetic and environmental variation, while  
38 rapid change in sex differences over time is more likely due to environmental change. One  
39 case of rapid change in sex differences is human lifespan, which has become increasingly  
40 female-biased in recent centuries. Long-term consequences of variation in the early-life  
41 environment may, in part, explain such variation in sex differences, but whether the early-life  
42 environment mediates sex differences in life-history traits is poorly understood in animals.  
43 Combining longitudinal data on 60 cohorts of pre-industrial Finns with environmental data,  
44 we show that the early-life environment is associated with sex differences in adult mortality  
45 and expected lifespan. Specifically, low infant survival rates and high rye yields in early-life are  
46 associated with female-bias in adult lifespan. These results support the hypothesis that  
47 environmental change has the potential to affect sex differences in life-history traits in natural  
48 populations of long-lived mammals.

49

## 50 **Introduction**

51 Phenotypic change in traits occurs over time through changing effects of both genetics and  
52 the environment. For example, lifespan has increased dramatically in humans over the last  
53 200 years in most industrialised nations (Beltrán-Sánchez *et al.* 2015) and, interestingly, this  
54 increase has been stronger in females. As a result, lifespan has become progressively more  
55 female-biased in many contemporary populations (Austad & Fischer 2016; Colchero *et al.*  
56 2016). Changes in sex differences are generally unlikely to occur rapidly through evolutionary  
57 processes because selection favouring sex differences frequently affects genes with similar  
58 effects in either sex, thus inducing intralocus sexual conflict (Lande 1979; Bonduriansky &  
59 Rowe 2005; Griffin *et al.* 2013). This suggests a role for environmental effects to mediate rapid  
60 shifts in sex differences in human lifespan, and traits in other natural populations. However,  
61 for life-history traits, the link between ecology and sex differences remains poorly understood,  
62 not just in humans, but across taxa (Brooks & Garratt 2017). Longitudinal series of data  
63 available on human populations (Stearns *et al.* 2010) can therefore prove to be a valuable  
64 asset when assessing the potential importance of environmental conditions for sex  
65 differences in life-history traits for natural populations of other long-lived animals, a critical  
66 issue given current rates of environmental change.

67 Particularly poor environmental conditions during development and early-life have been  
68 linked to long-lasting adverse effects on health, reproductive success, and mortality rates in a  
69 range of species (Lindström 1999; Metcalfe & Monaghan 2001), including humans (Lummaa  
70 & Clutton-Brock 2002). However, whether such factors are associated with variation in sex  
71 differences in human life-history traits remains unexplored (Austad 2006; Austad & Fischer  
72 2016). While evidence that changing patterns of sex differences in key life-history traits can

73 result from variation in early-life environments is very limited, two recent studies, in a bird  
74 (Wilkin & Sheldon 2009) and a mammal (Garratt *et al.* 2015), do provide some support to this  
75 hypothesis. In Great Tits (*Parus major*), variation in altitude and size of the natal territory have  
76 been linked to recruitment and lifespan respectively in males, but not females, suggesting that  
77 variation in the natal environment can drive sex differences in life history traits. In roe deer  
78 (*Capreolus capreolus*), higher juvenile mortality rates, which are indicative of a less favorable  
79 environment, were associated with higher adult lifespan in females and lower adult lifespan  
80 in males. As a result, lifespan is expected to become more female-biased in more harsh early-  
81 life conditions (Garratt *et al.* 2015). Harsh early-life conditions are either proposed to better  
82 prepare individuals for similar conditions later in life (predictive adaptive response, PAR, see  
83 Gluckman & Hanson 2004), or reduce adult lifespan because of lasting damage to the  
84 individual, as suggested by the silver-spoon hypothesis (Wilkin & Sheldon 2009). Alternatively,  
85 weaker individuals may be less likely to survive to adulthood when conditions are harsh,  
86 leaving only the most viable individuals within the adult population and thus increasing  
87 average lifespan of the cohort (viability selection, see Garratt *et al.* 2015).

88 We use longitudinal data from individuals born across a 60-year period in a pre-industrial  
89 Finnish population combined with detailed data on year-to-year environmental conditions to  
90 test for associations between sex differences in life-history traits (adult mortality and adult  
91 life expectancy) and environmental variation. Sex differences in adult lifespan may have a  
92 number of underlying causes, such as sex differences in the tendency to access and use  
93 medical care (Bertakis *et al.* 2000; Vaidya *et al.* 2012), which may also be associated with  
94 industrialisation. By using data from a pre-industrial population (Bolund *et al.* 2015) with no  
95 access to such advances, we are able to assess the ecology of sex differences in adult mortality  
96 linked to variation in early-life conditions, across cohorts with variable magnitude of sex

97 differences, while minimizing such confounding effects. During the study period (1791-1850)  
98 the Finnish population depended on small-scale farming of rye and barley for their livelihood,  
99 supplemented by fishing and hunting. Unpredictable climatic conditions led to frequent crop  
100 failures and, subsequently, famines and disease outbreaks were common (Hayward *et al.*  
101 2013, 2016). Environmental variation occurred throughout the study period, which we first  
102 quantify as variation in the infant survival rate, as previously used in a similar study in roe deer  
103 (Garratt *et al.* 2015) and other species (e.g. Robinson *et al.* 2008; Martin & Festa-Bianchet  
104 2011), as a broad-scale indicator of environmental severity. We then extend the study to a  
105 more specific ecological context, utilising data on annual crop yields and climate in a subset of  
106 cohorts for which this data is available.

107

## 108 **Material and Methods**

### 109 *Human data*

110 The Lutheran Church in Finland has been required to record births, deaths, and marriages  
111 since the 1700s and we use these data for 1,988 individuals born in the parish of Ikaalinen in  
112 south-west Finland between 1791 and 1850 (60 birth cohorts). The number of individuals born  
113 per cohort over the 60 years averaged  $67.2 \pm 23.96$ , ranging between 28 and 114, with 49.3%  
114 of those born surviving to adulthood. The individuals included in our study sample survived to  
115 adulthood (age 15), did not migrate before adulthood, and had known socio-economic status  
116 during childhood (socio-economic status of the father). Childhood socio-economic status may  
117 affect the relationship between adult mortality risk and the environment by determining  
118 susceptibility to environmental factors. We therefore estimate and compare the sex-specific  
119 responses to environmental variation while controlling for childhood socio-economic status.

### 120 *Environmental data*

121 We explored the environment which individuals were exposed to in childhood taken as both  
122 *i)* a three-year mean around the birth year, combining gestation and the early post-natal  
123 period, and *ii)* the average environment over the first ten years of life (juvenile period).  
124 Throughout the study, infant survival rates were calculated per cohort as the proportion of  
125 individuals surviving to the age of five years such that; for estimates at the birth year, it is the  
126 proportion of individuals born in the focal year, the preceding year, and the following year  
127 that survive to age five; for the 10-year juvenile period estimates, it is the proportion of  
128 individuals born in the focal year, or the following 9 years, that survive to age 5. The annual  
129 infant survival rates were highly variable among the 60 cohorts, ranging between 50% and  
130 96% survival to the age of five and at an average of 73% survival.

131 We use data on crop sowing success (yields) as indicators of local grain availability and  
132 nutritional stress. Crop yield is measured as the ratio of grain sown to grain harvested by  
133 weight, and the annual success varied between 2.71 and 10.47 for barley, and 2.81 and 11.06  
134 for rye, with mean yields of 6.07 and 6.27 for barley and rye respectively. Crop yields, while  
135 not measuring the *per capita* availability of crops, indicate the between-year variation in  
136 harvest success without being biased by factors such as cultivation area or intensity, providing  
137 a relative measure of nutritional stress (Hayward *et al.* 2012). Data on barley yield was  
138 available to produce 3-year estimates for 1543 individuals across 37 cohorts (mean cohort size  
139 = 41.7), and 10-year estimates for 1345 individuals across 30 cohorts (mean cohort size = 44.8).  
140 Data on rye yield was available for 3-year estimates for 1588 individuals across 40 cohorts  
141 (mean cohort size = 39.7) and for 10-year estimates for 1345 individuals across 30 cohorts  
142 (mean cohort size = 44.8). We also use data on spring temperature to assess the effect of  
143 variation in climatic conditions. Reconstructed annual spring temperature data for south-west  
144 Finland (Holopainen *et al.* 2009) showed a mean temperature of 2.56°C and range of 0.28°C  
145 to 7.19°C for the study period. Data on spring temperature was available for 3-year estimates  
146 for 1762 individuals across 46 cohorts (mean cohort size = 38.3), and for 10-year estimates for  
147 1789 individuals across 47 cohorts (mean cohort size = 38.1).

148 Crop yields and spring temperatures are known to have affected life-history traits in the study  
149 population (Rickard *et al.* 2010; Hayward *et al.* 2013): a great famine during the late-1860s  
150 largely resulted from failed crops following a late onset of spring bringing unusually high infant  
151 mortality; in 1867, 48.0% of children born in Ikaalinen died before the age of five. There were  
152 no secular trends (Figure S1) for annual measures of infant survival (generalized linear model:  
153  $\beta = 0.001 \pm 0.017$ ,  $p = 0.951$ ), barley yield (linear model [lm]:  $\beta = -0.004 \pm 0.023$ ,  $p = 0.851$ ), rye

154 yield (lm:  $\beta = -0.011 \pm 0.021$ ,  $p = 0.589$ ), and spring temperature (lm:  $\beta = -0.010 \pm 0.015$ ,  $p =$   
155 0.481).

### 156 *Statistical analysis*

157 Effects of the environmental factors on sex differences in adult mortality were assessed in Cox  
158 proportional hazards models using the 'coxph' function in the 'survival' R package (Therneau  
159 2012), using the individuals that survived to adulthood as the study population. Individual  
160 models were constructed for each environmental factor, with lifespan as the response  
161 variable, and including fixed effects of sex (two-level factor), childhood socio-economic status  
162 (two-level factor; 'land-owners' and 'landless'; see Bolund et al 2015), and the environmental  
163 factor of interest, as well as their interactions (sex x environment x socio-economic status, sex  
164 x environment, socioeconomic status x environment) and the random effect of cohort. These  
165 models estimate values of the sex-specific regression coefficients ( $\beta$ ), herein denoted by  
166 subscript M (male) and F (female) as appropriate. Cox proportional hazards models were used  
167 as they also allow the inclusion of individuals with "censored" lifespan, i.e. those with a  
168 minimum known lifespan inferred from records of other events ( $n = 663$ ) such as church  
169 attendance, child birth, marriage, and migration, rather than limiting analyses to a potentially  
170 biased sample of individuals with a known lifespan. We did not control for: mother identity,  
171 because only a small number of offspring per mother survive to adulthood (and for a large  
172 proportion only 1) and thus mother identity would largely explain the same variation as  
173 individual identity; maternal mortality during childhood, because this information was not  
174 available for all our study subjects and the negative effects of maternal death on individual  
175 mortality risk appear to be short-term in the study population, with effects present in  
176 childhood (Lahdenperä *et al.* 2011); and reproductive status of the individual (or number of

177 children), as this may have been inherently biased by survival because short-lived individuals  
178 had less opportunity to reproduce. We did not test models that contain multiple  
179 environmental factors simultaneously (e.g. barley and rye yield) and their interactions  
180 because; *i*) there were strong correlations among the fixed environmental effects measured  
181 over the 10-year juvenile period ( $r_{\text{Barley, Rye}} = 0.69$ ,  $p < 0.001$ ;  $r_{\text{Barley, Temperature}} = -0.55$ ,  $p = 0.001$ ;  
182  $r_{\text{Rye, Temperature}} = -0.44$ ,  $p = 0.013$ ), and *ii*) not all individuals included in the analyses had data for  
183 all factors for the same years, thus doing so reduces sample size and power.

184 Models were reduced by backwards elimination, retaining terms which substantially improved  
185 model fit (as judged by AIC, Table S2), the two-way interaction of interest in this study (sex ×  
186 environmental factor), and any main effects for factors retained in higher order terms. The  
187 assumption of proportional hazards was checked using the 'cox.zph' function and all models  
188 considered in the analyses satisfy this assumption. From the model summaries, we extracted  
189 estimates of the sex-specific coefficients, and differences in sex-specific coefficients, along  
190 with estimates of the standard error and p values for a null hypothesis that the coefficient, or  
191 difference between coefficients, is zero. Predictions of sex-specific adult life expectancy in  
192 relation to infant survival rate and environmental factors were also drawn from these reduced  
193 models, using the 'survfit' function in the 'survival' R package, while removing the random  
194 effect of cohort as survfit cannot handle models containing random effects terms. Two further  
195 Cox proportional hazards models were fitted with fixed effects of only cohort, and both sex  
196 and cohort, and compared by likelihood ratio testing to test for variation in sex-specific adult  
197 mortality risk. Population-wide predicted life expectancy at adulthood was also estimated  
198 using the latter model and survfit function.

199 **Results**

200 Within this study population, there was considerable variation in sex differences in adult  
201 mortality ( $X^2 = 17.73$ ,  $p < 0.001$ ) illustrated by clear sex differences in survival curves (Figure  
202 1A). As a result, adult life expectancy was predicted to be female-biased, to a varying degree,  
203 in all cohorts (Figure 1B), with expected lifespan at adulthood on average being 62.30 years in  
204 males and 66.41 years in females.

205 *Sex-specific responses to early-life environmental variation*

206 Cohorts which were exposed to lower rates of infant survival at the time of birth (3-year)  
207 expressed increased adult mortality risk in males ( $\beta_M = -1.128 \pm 0.562$ ,  $p = 0.045$ ), consistent  
208 with the silver-spoon hypothesis, while no association was present in females ( $\beta_F = 0.132 \pm$   
209  $0.587$ ,  $p = 0.820$ ). When considering the early-life environment over the 10-year juvenile  
210 period, males show a qualitatively similar but not statistically significant association with  
211 infant survival rate ( $\beta_M = -1.148 \pm 0.927$ ,  $p = 0.215$ ), while a stronger association closer to  
212 statistical significance was found in females, with a lower adult mortality risk in cohorts  
213 exposed to low infant survival rates during early-life ( $\beta_F = 1.890 \pm 0.989$ ,  $p = 0.056$ , Figure 2A),  
214 consistent with the viability selection hypothesis. While averaging the environment over a  
215 longer period of time will reduce the environmental variation between cohorts and potentially  
216 weaken the ability to detect effects early in childhood, the developmental period in humans  
217 is extraordinarily long compared to other long-lived mammals, thus the longer time period  
218 allows us to test for any effects which accumulate through later phases of development which  
219 would be missed in analysing just the birth period. Because our results illustrate that the  
220 environment experienced during the entire early-life period is important in determining adult

221 mortality risk, we herein focus on effects over the 10-year juvenile period for brevity (but see  
222 Table S1 for the corresponding results for environmental conditions around the time of birth).  
223 Cohorts exposed to better nutritional conditions during the juvenile period, as measured by  
224 higher yields of barley, have lower adult mortality risk in females ( $\beta_F = -0.272 \pm 0.111$ ,  $p =$   
225  $0.015$ ) but not males ( $\beta_M = -0.037 \pm 0.107$ ,  $p = 0.733$ ). Inspection of survival curves (Figure 3A)  
226 shows that adult mortality is decreased for females when yields of barley are high in early-life.  
227 Similarly, exposure to high yields of rye during the juvenile period was also associated with  
228 reduced adult mortality in females ( $\beta_F = -0.187 \pm 0.065$ ,  $p = 0.004$ ) but not males ( $\beta_M = 0.038$   
229  $\pm 0.056$ ,  $p = 0.501$ , Figure 3C). These results support that the nutritional environment during  
230 early-life has the potential to affect adult mortality and life expectancy, such that nutritionally  
231 rich environments were associated with longer life expectancy in females, a result consistent  
232 with the silver-spoon hypothesis, but not males (Figure 3B & 3D). Spring temperature during  
233 the juvenile period was positively associated with adult mortality in males ( $\beta_M = 0.107 \pm 0.052$ ,  
234  $p = 0.039$ ) but not females ( $\beta_F = 0.068 \pm 0.052$ ,  $p = 0.192$ ), such that males from cohorts which  
235 endured colder conditions during the juvenile period had lower adult mortality risk (Figure  
236 4A). This supports that variation in adult life expectancy may be mediated by climatic variation  
237 in early-life, with expected lifespan decreasing in males more than females as spring  
238 temperatures during the juvenile period increase (Figure 4B).

### 239 *Sex differences in responses to early-life environmental variation*

240 Comparison of sex-specific coefficients in response to infant survival rates during the juvenile  
241 period suggests that sex differences in adult mortality risk were associated with variation in  
242 the early-life environment ( $\beta_F - \beta_M = 3.039 \pm 1.326$ ,  $p = 0.022$ ). Survival curves suggest that sex  
243 differences in mortality were smaller in cohorts exposed to high rates of infant survival (Figure

244 2A). As a result, adult life expectancy was female-biased in cohorts exposed to low rates of  
245 infant survival during the juvenile period (Figure 2B). For all three ecological factors (barley  
246 yield, rye yield, and spring temperature), we find statistically significant responses in one sex  
247 and smaller non-detectable responses in the other, supporting that various aspects of the  
248 environment experienced during the juvenile period simultaneously contributed to sex  
249 differences in adult mortality risk. Formal comparison of the sex-specific estimates suggests  
250 that adult mortality risk in males and females showed different responses to rye yields ( $\beta_F - \beta_M$   
251 =  $-0.224 \pm 0.084$ ,  $p = 0.008$ ). However, we find no statistically significant differences between  
252 male and female specific associations with barley yields ( $\beta_F - \beta_M = -0.235 \pm 0.151$ ,  $p = 0.120$ ) or  
253 spring temperature ( $\beta_F - \beta_M = -0.039 \pm 0.072$ ,  $p = 0.591$ ) experienced during the juvenile period.  
254 Overall, these results support that variation in early-life environment can have lasting  
255 consequences mediating variation in sex differences in life-history traits.

256

## 257 **Discussion**

258 The early-life environment is known to have lasting effects on life history traits (Hayward et  
259 al. 2013), but whether variation in the early-life environment is linked to variation in the extent  
260 and direction of sex differences in such traits is poorly understood in natural animal  
261 populations. Our study supports the conjecture that the environment experienced during  
262 early-life can have long-lasting consequences for sex differences in life-history traits such as  
263 adult mortality risk and expected lifespan, highlighting the importance of determining the  
264 ecological drivers of sex differences both in humans and other natural animal populations,  
265 particularly in this era of rapid environmental change (Simmonds & Coulson 2015). The recent  
266 transitions in human mortality and fertility patterns as a consequence of rapid environmental  
267 change are among the most striking described for any species (Scranton *et al.* 2016), and the  
268 longitudinal records available on humans across varying environmental conditions make them  
269 a valuable system which can be used to study the effects of environment on sex differences  
270 in life-history traits in a general context which is informative across taxa.

271 Using infant survival rate as an indicator of environmental severity, as in previous ecological  
272 studies, we show that harsh early-life environments were associated with increased sex  
273 differences in mortality risk and female-bias in adult life expectancy. This result is concordant  
274 with a recent study in roe deer (Garratt *et al.* 2015), but not with modern human trends which  
275 are characterized by increasing infant survival and female-bias in life expectancy. While infant  
276 survival rates are used within ecology as a proxy for the severity of environmental conditions,  
277 they can be simultaneously affected by a wide range of ecological factors including nutritional  
278 and abiotic stressors, as well as epidemics of infectious disease which were common at the  
279 time of our study population (Hayward *et al.* 2016) and were not incorporated into our study.

280 It is likely that the factors determining infant survival rates differ greatly between our study  
281 population and current human populations, which could offer an explanation for the  
282 discrepancy between our results and modern trends. Our results from three environmental  
283 factors (barley yields, rye yields, and spring temperature), which more directly assess potential  
284 environmentally induced stress, suggest that less harsh environments, those with better food  
285 availability and warmer temperatures, led to further male-bias in adult mortality risk and  
286 female-bias in expected lifespan. This latter result can be considered as being consistent with  
287 trends over recent generations of humans. Taken together, these results highlight that  
288 conclusions drawn using coarse demographic measures of the environment, such as infant  
289 survival rates which are affected simultaneously by many different factors such as disease and  
290 more stochastic events, may differ to those using more direct measures of the environment  
291 due to the complex nature of coarse measures.

292 Sex differences in life-history traits have been linked to early-life environmental variation in a  
293 limited number of studies in natural environments, in a mammal (Garratt *et al.* 2015) and a  
294 bird (Wilkin & Sheldon 2009) species. While such links have not previously been demonstrated  
295 in human populations, knowledge of the general importance of early-life effects in  
296 determining sex differences in life-history traits is certainly also lacking for natural populations  
297 (Brooks & Garratt 2017). Further female-bias in adult life expectancy was predicted to occur  
298 in our study population when there were good crop yields or warmer spring temperatures,  
299 with these occurring through decreasing female mortality in the case of crops, and increasing  
300 male mortality in the case of spring temperature. The sex-specific associations with rye yield,  
301 and the general sex-specific pattern found across environmental factors, suggest that  
302 improved access to food and reduced environmental stress lead to increased sex differences  
303 in adult mortality risk, and increased female-bias in life expectancy. In the birth-year centred

304 results on crop yields, models featuring the three-way interaction between sex, environment,  
305 and childhood socio-economic status, further support that the environment and susceptibility  
306 to the environment, which may be mediated by socio-economic status, can affect sex  
307 differences in life history traits (Figure S3). While human populations in most industrialised  
308 nations have moved far beyond the levels of nutritional and climate induced stress that the  
309 population in our study would have experienced, our results nonetheless give the first  
310 evidence that sex-specific responses to the early-life environment may contribute to sex  
311 differences in life-history traits in humans. Indeed, ecological variation during early  
312 development even in contemporary developed countries has been linked to a range of health  
313 outcomes in adulthood (Huxley *et al.* 2007; Mandl *et al.* 2009; Belbasis *et al.* 2016), showing  
314 that between-individual differences in developmental conditions continue to exert long-term  
315 effects in humans. Furthermore, these results make a valuable contribution to our broader  
316 understanding of the link between variation in the early-life environment and sex differences  
317 in life-history traits.

318 Sex differences in life expectancy and mortality could be linked to the environment if lifespan  
319 responds differently to environmental effects in males and females. At the individual level,  
320 harsh environments may either increase (PAR) or decrease (silver-spoon) expected lifespan,  
321 while cohorts may also bear the signature of viability selection. In the latter case, weaker  
322 individuals that would be short-lived as adults are more likely to die prior to adulthood when  
323 exposed to harsh conditions and, if environmentally determined rates of viability selection are  
324 sex-specific, sex-bias in pre-adulthood mortality would be associated with sex differences in  
325 adult lifespan. Males typically invest more than females in the development of costly sexually  
326 selected traits (Vinogradov 1998; Bonduriansky *et al.* 2008), such as large body size, at the  
327 expense of reduced survival probability. Under harsh environmental conditions, expensive

328 male growth requirements could thus lead to male-biased pre-adult mortality, thereby  
329 increasing viability selection in males and increasing male lifespan more so than in females in  
330 the resulting cohort. Our data indicates that this may be the case in humans, as female-bias in  
331 expected lifespan was lower when pre-adulthood mortality was male-biased, such that sex-  
332 biased rates of infant mortality are associated with the sex differences in adult life expectancy  
333 (Figure S4). Disentangling the mechanisms underlying sex differences in life-history traits is a  
334 substantial challenge which is yet to be addressed, and while our results indicate that sex-  
335 differences in viability selection may be a factor here, there is a pressing need to address this  
336 issue more thoroughly both in humans and other natural systems. Our results from the Finnish  
337 pre-industrial population offer support to the hypothesis that environmental changes can help  
338 to explain increasing sex differences in life-history traits such as lifespan, mortality, and aging,  
339 as has been increasingly common in human populations since industrialisation. In a broader  
340 context, our results suggest that accelerating environmental change could dramatically alter  
341 patterns of sex differences in life-history traits in natural populations of other animals.

342

343 **Acknowledgments**

344 The authors thank the Academy of Finland (VL, RMG), University of Stirling Research  
345 Fellowship (ADH), the Swedish Research Council (EB, AAM), and the European Research  
346 Council (AAM) for funding.

347

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466 **Figure Legends**

467 **Figure 1. Survival and sex differences in expected adult lifespan.** A) Survival curves for the  
468 1988 individuals used in the study (see methods), with males (blue) showing lower survival  
469 than females (red). B) Sex-bias (SB;  $1-[M/F]$ ) in life expectancy at adulthood, as predicted by  
470 the model, was female-biased for all cohorts.

471 **Figure 2. The relationship between 10-year infant survival rates and adult life-history.** A)  
472 Females (red) showed consistently higher rates of survival than males (blue), and this  
473 difference appears to be smaller in cohorts with above average 10-year infant survival rate  
474 (solid lines) than in cohorts with infant survival rates below average (dashed lines). B) Life  
475 expectancy at adulthood was expected to be higher in females (red) than males (blue) when  
476 infant survival rates were low, and for males to outlive females when infant survival rates were  
477 high, though 95% confidence intervals (shaded areas) overlap for much of this area. Sample  
478 size of 1988 individuals across 60 cohorts.

479 **Figure 3. The relationship between 10-year crop yields and adult life-history.** A) Survival in  
480 males (blue) in cohorts exposed to below average barley yields (dashed lines) were similar to  
481 those exposed to above average yields of barley (solid lines), while female (red) survival was  
482 increased when yields of barley were high. B) Life expectancy at adulthood was more strongly  
483 associated with barley yield in females (red) than males (blue), such that lifespan was  
484 predicted to become female-biased when barley yields were high, and male-biased when  
485 barley yields were low. Confidence intervals do not show for higher yields of barley because  
486 few cohorts experience high yields (Figure S1). C) Similarly to barley yield, rye yield showed  
487 greater effects in females than males, with survival increasing in females when rye yields were

488 high, leading to D) predictions of increasing female-biased lifespan when rye yields were high.

489 Sample size of 1345 individuals across 30 cohorts.

490 **Figure 4. The relationship between 10-year spring temperature (°C) and adult life-history.** A)

491 Survival was reduced in cohorts exposed to above average spring temperatures (solid lines) in

492 both males (blue) and females (red) relative to cohorts exposed to below average spring

493 temperatures. B) Adult life expectancy showed mild declines in both sexes in cohorts exposed

494 to higher spring temperatures, but predicted increasing female-bias in lifespan as temperature

495 increases. Sample size of 1789 individuals across 47 cohorts.

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