

**Social Sciences: Social Sciences**

**EARLY-LIFE DISEASE EXPOSURE AND ASSOCIATIONS WITH ADULT SURVIVAL, CAUSE OF DEATH AND REPRODUCTIVE SUCCESS IN PRE-INDUSTRIAL HUMANS**

**Short title: Early-life disease and later fitness in humans**

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1   **ABSTRACT**

2

3   A leading hypothesis proposes that increased human lifespan since 1850 has resulted from  
4   decreased exposure to childhood infections, which has reduced chronic inflammation and  
5   later-life mortality rates, particularly from cardiovascular disease, stroke and cancer. Early-  
6   life cohort mortality rate often predicts later-life survival in humans, but such associations  
7   could arise from factors other than disease exposure. Additionally, the impact of early-life  
8   disease exposure on reproduction remains unknown, and thus previous work ignores a major  
9   component of fitness, through which selection acts upon life-history strategy. We collected  
10   data from seven 18<sup>th</sup> and 19<sup>th</sup> century Finnish populations experiencing naturally-varying  
11   mortality and fertility levels. We quantified early-life disease exposure as the de-trended  
12   child mortality rate from infectious diseases during an individual's first five years, controlling  
13   for important social factors. We found no support for an association between early-life  
14   disease exposure and all-cause mortality risk after age 15 or 50. We also found no link  
15   between early-life disease exposure and probability of death specifically from cardiovascular  
16   disease, stroke or cancer. Independently of survival, there was no evidence to support  
17   associations between early-life disease exposure and any of several aspects of reproductive  
18   performance, including lifetime reproductive success and age at first birth, in either males or  
19   females. Our results do not support the prevailing assertion that exposure to infectious  
20   diseases in early life has long-lasting associations with later-life all-cause mortality risk or  
21   mortality putatively linked to chronic inflammation. Variation in adulthood conditions could  
22   therefore be the most likely source of recent increases in adult lifespan.

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26 **SIGNIFICANCE STATEMENT**

27

28 Why has human life expectancy increased since 1850? A leading hypothesis proposes that  
29 limited exposure to childhood infections has reduced lifelong inflammation and enhanced  
30 survival, but tests of this hypothesis typically use all-cause mortality rates to estimate disease  
31 exposure. Meanwhile, links between early-life disease and reproduction have been neglected.

32 We used data from pre-industrial Finnish populations to show that early-life disease exposure  
33 was not associated with all-cause mortality, mortality from cardiovascular disease, stroke and  
34 cancer, or reproductive success. Our study therefore does not support the prevailing  
35 contention that reduced exposure to early-life infections has increased life expectancy in  
36 modern populations.

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51    **INTRODUCTION**

52

53    In industrialized nations, human life expectancy at birth has increased by around three  
54    months per year between 1840-2002 (1). While this is linked to improvements in child  
55    survival rates, there have been considerable increases in life expectancy of adults: in the UK  
56    in 1843, the mean expected age at death of a 20-year old was 60, while in 2011 it was 81 (2,  
57    3), a 50% increase in remaining lifespan. Longer adult life expectancy has been accompanied  
58    by increased incidence of previously rare chronic health problems and thus identifying the  
59    drivers of enhanced lifespan is an important question for global health. Declines in adult  
60    mortality rate since the 18<sup>th</sup> century in Europe were more closely linked to year of birth than  
61    to the year in which mortality was assessed, leading to the hypothesis that: "...the expectation  
62    of life was determined by the conditions which existed during the child's earlier years" (4).  
63    The 'cohort morbidity phenotype' hypothesis suggests that the link between declining birth  
64    year mortality rate and increasing adult survival is due to reduced exposure to infectious  
65    diseases in early life and reduced lifelong burden of chronic inflammation (5–8). Thus,  
66    individuals experiencing exposure to infectious disease in early life are expected to have  
67    higher adulthood mortality risk.

68

69    It is hypothesized that early-life exposure to infectious diseases could increase later-life  
70    mortality risk through a number of physiological mechanisms (6). Bacterial and viral  
71    infections elicit inflammatory immune responses, and chronic inflammation is associated  
72    with atherosclerosis, increased risk of stroke, cardiovascular disease (CVD) and mortality (9).  
73    Thus, early-life disease exposure is expected to be linked to adult mortality risk from these  
74    causes in particular (6). Support for this hypothesis comes from studies showing that  
75    exposure to high levels of early-life cohort mortality or infection is associated with middle-

76 age mortality from CVD (10, 11); that exposure to infection is associated with chronically  
77 high levels of inflammation (12, 13); that high levels of inflammation are associated with  
78 CVD (14). This physiological mechanism is underpinned by the fact that natural selection  
79 should favour the expression of robust immune responses at young ages, even at the cost of  
80 lower later-life survival (15).

81

82 Building on such knowledge, studies have tested the ‘cohort morbidity phenotype’ hypothesis  
83 by showing that early-life infections are associated with chronic inflammation and impaired  
84 later health (16, 17). A drawback is that these studies are performed on modern populations  
85 with good health care and therefore cannot estimate associations between early-life disease  
86 exposure and natural mortality risk. A second approach is to use historical data on natural-  
87 mortality pre-industrial populations and to link early-life cohort mortality (a proxy for early-  
88 life disease exposure) with later-life mortality (18–20). Studies of historical populations have  
89 however experienced at least one of several shortcomings. First, many studies are based on  
90 population-level statistics, meaning they cannot identify factors allowing some individuals to  
91 live long lives, despite being born in high-mortality cohorts. Second, most studies only  
92 consider all-cause cohort mortality rate as a measure of disease exposure, which may be  
93 driven by nutrition and social factors as well as disease. Third, it is often unknown which  
94 types of deaths are associated with early stressors, obfuscating the mechanism through which  
95 early-life disease exposure is associated with later mortality.

96

97 The majority of studies of early-life disease exposure have focused solely on associations  
98 with later-life mortality, despite the fact that selection on life-history strategy should act  
99 through both survival and reproductive success. Life-history theory predicts that high  
100 mortality risk in early life should lead to earlier age at maturity (21), but associations between

101 early-life disease exposure and between-individual variation in reproductive success have  
102 been neglected. A series of studies have shown that early life trauma, such as sibling  
103 mortality, parental separation, disadvantageous birth location and famine (22–25) are  
104 associated with earlier age at first reproduction and/or larger family size, suggesting adoption  
105 of a ‘fast’ life-history in high-stress, high-mortality environments. In contrast, childhood  
106 infections may have a non-lethal but damaging impact upon physiology which may impede  
107 reproduction during adulthood (26). These could include infections after birth leading to  
108 stunted growth in childhood and eventually reduced adult size (27), or even infections of  
109 focal individuals’ mothers leading to reduced birth weight, which may lead to reduced birth  
110 weight of focal individuals’ offspring (28, 29). Nevertheless, it has been shown that the  
111 direction of the association between final adult height and fertility is variable across  
112 populations (30).

113

114 We used church register data collected from seven Finnish populations across the 18<sup>th</sup> and  
115 19<sup>th</sup> centuries to determine the association between early-life exposure to infectious diseases  
116 and (1) later-life mortality risk; (2) cause of death; (3) reproductive success. These  
117 populations had no access to modern contraception or healthcare (31) and as a result ~40% of  
118 children died of disease before age 15. Having survived to age 15, mean lifespan in our  
119 sample was ~56 years and individuals produced an average of ~4 surviving offspring, with  
120 >99% of individuals born to parents aged <50. We used cause-of-death data recorded by  
121 churches to adopt a measure of disease exposure based on the local probability of child death  
122 from infectious diseases during the early years of an individual’s life. Crucially, we  
123 controlled for other salient variables reflecting early-life conditions and known to influence  
124 fitness in these populations (SI Appendix, Table S1 and Fig. S1), including social class at  
125 birth (32), twin status (33), inter-birth interval (34) and season of birth (35).

126

127 **RESULTS**

128

129 There was considerable between-year variation in the prevalence of deaths from infectious  
130 diseases in the early years of our study individuals' lives. We used these data to calculate a  
131 measure of early-life exposure to infectious diseases for each of the first five years of life of  
132 each individual, based on the proportion of children dying from infectious diseases in their  
133 parish in each year. We recorded the number of children aged 0-10 alive in each parish in  
134 each year, 1751-1855 (a total of 195,170 person-years). In 194,182 (>99.5% of) cases, the  
135 child had a known cause of death or survived, while in the remaining 988 cases, the child  
136 died of an unknown cause; these 988 cases were not included in our calculations of early-life  
137 disease exposure (SI Appendix, Table S2). During these years, 4780 child deaths of known  
138 cause were recorded, with 73% due to infectious diseases (SI Appendix, Table S3). We  
139 calculated annual parish-specific variation in child mortality rate from infectious diseases by  
140 dividing the number of child (aged 0-10) deaths in each parish due to infections in each year,  
141 1751-1855, by the number of children alive in each parish in the same year. The child  
142 mortality rate from infectious diseases varied substantially across the study period in all  
143 seven parishes (Fig. 1). We de-trended the child mortality rate from infectious diseases by  
144 applying a Hodrick-Prescott filter to data from each parish using the R package 'mFilter',  
145 with a filtering term of 100. This decomposed variation in child mortality rate from infectious  
146 diseases in each parish into a long-term smoothed 'trend' component (Fig. 1) and the short-  
147 term deviation from the trend, or the 'cycle' (Fig. 2). The 'cycle' component for each year  
148 was multiplied by 100 to aid model convergence, and we refer to this measure herein as  
149 annual 'disease exposure' (Fig. 2). Each individual was therefore defined by their potential  
150 early exposure to infections, rather than simply the level of all-cause cohort mortality (8, 18,

151 19), on a continuous quantitative scale, accounting for temporal changes in conditions, and  
152 specific to the parish in which they were born. We designed analyses in order to determine  
153 whether our measure of disease exposure in each of the first five years of life was associated  
154 with later-life survival and reproductive success.

155

## 156 **Mortality risk**

157

158 Our mixed-effects Cox model analyses showed that greater disease exposure in each of the  
159 first five years of life was not associated with increased mortality risk in adulthood (N=7283).  
160 Controlling for a higher mortality risk among poor individuals and males, and variation  
161 between parishes, as well as other fixed and random effects (Table 1), disease exposure  
162 measured in each of the first five years of life was not significantly associated with hazard of  
163 mortality after age 15 and nor was the association between early-life disease exposure and  
164 mortality risk dependent upon parish, social class, or sex (SI Appendix, Table S4). When we  
165 restricted our analyses to individuals who survived to at least the age of 50 (N=3822) and  
166 who were therefore most likely to experience stroke, CVD and cancer, we again found no  
167 evidence for an association between early-life disease exposure and mortality risk and no  
168 indication of interactions with individual parish, social class or sex (SI Appendix, Table S4).

169

## 170 **Cause of death**

171

172 Mixed-effects Cox models showed that there was no significant association between early-  
173 life disease exposure and cause of death. We tested for an association between disease  
174 exposure in each of the first five years of life and death from CVD, stroke and cancer, on  
175 which the ‘cohort morbidity phenotype’ focuses (8). Individuals were scored as 1 (a ‘failure’)

176 if they died of CVD, stroke or cancer, and 0 if they died of other causes or were censored  
177 without dying (SI Appendix, Table S5 & Fig. S2). As above, we first performed the analysis  
178 on all individuals who survived to at least the age of 15. There was no statistical support for  
179 an association between early-life disease exposure in any of the first five years of life and  
180 death from CVD, stroke or cancer (SI Appendix, Table S6). Although there was some  
181 evidence to support an interaction between early disease exposure and social class (Disease  
182 Year 4:Social class (Poor) Hazard = 0.79; 95% C.I.=0.62-1.00;  $\chi^2_1=4.17$ ; P=0.0.041)  
183 suggesting that Poor individuals with higher early-life disease exposure in the fourth year of  
184 life were less likely to later die of infection (Table S7), this interaction was only weakly  
185 statistically supported over the base model with  $\Delta AIC=-2.02$  (36). Nevertheless, it suggested  
186 a trend for a mildly protective effect of early disease exposure and later death risk from CVD,  
187 stroke or cancer in Poor individuals. We also repeated this analysis on individuals who  
188 survived to at least age 50, since the majority (263/346) of individuals affected by these death  
189 causes were aged over 50. Once again, there was no evidence that early-life disease exposure  
190 was associated with cause of death, nor any evidence for interactions between disease  
191 exposure and parish, social class or sex (SI Appendix, Table S6).

192

### 193 **Lifetime reproductive success**

194

195 We next assessed links between early-life disease exposure and later-life reproductive  
196 success in individuals who married and survived to at least age 50 and who had therefore  
197 completed their reproductive lifespans with the opportunity to reproduce. We analysed data  
198 from females and males separately, due to their differing social and physiological factors  
199 affecting reproductive schedules. First, we used linear mixed-effects models to show that  
200 there was no association between early-life disease exposure and Box-Cox transformed

201 lifetime reproductive success (LRS) in females (N=1512), defined as the number of offspring  
202 an individual produced which survived to age 15 (best-fitting single year, Disease year 1  
203 estimate =  $-0.0202 \pm 0.0112$ ;  $\chi^2_{(1)}=3.25$ ; P=0.071; SI Appendix, Table S8). This model  
204 controlled for variation between parishes ( $\chi^2_{(6)}=68.60$ ; P<0.001), lower LRS in Poor  
205 compared to Rich individuals (estimate=-0.24±0.04;  $\chi^2_1=44.77$ ; P<0.001) and non-significant  
206 effects of birth interval, birth season and twinning status. There was also no statistical support  
207 for interactions between early disease exposure and social class or parish (SI Appendix, Table  
208 S8). In males (N=1333), there was once again no support for an association between early-  
209 life disease exposure and LRS (best-fitting single year, Disease year 4  
210 estimate= $0.0068 \pm 0.0118$ ;  $\chi^2_1=0.33$ ; P=0.563), nor any support for interactions with social  
211 class or parish (SI Appendix, Table S8).

212

213 We next determined whether there were any links between early-life disease exposure and  
214 other components of reproduction which constitute LRS, namely age at first reproduction, the  
215 number of children born, or child survival rate. In neither sex, however, did we find any  
216 support for associations between early-life disease exposure and any of these measures of  
217 reproductive success, either as a main effect or in interactions with social class or parish (SI  
218 Appendix, Tables S9-S11).

219

## 220 **DISCUSSION**

221

222 It has been hypothesized that recent increases in human lifespan have been brought about by  
223 reduced exposure to infectious diseases in early life and reduction of lifelong inflammation  
224 and associated chronic diseases such as CVD, stroke and cancer (5–8). Tests of this  
225 hypothesis have used data from historical populations to support a link between cohort all-

226 cause mortality rate and later-life survival, but have not assessed later-life survival in relation  
227 to measures of early-life disease exposure based on disease prevalence, or determined how  
228 early disease exposure is linked to mortality risk specifically from CVD, stroke and cancer.  
229 Studies of how early-life disease exposure may impact upon later-life reproductive success  
230 are also extremely rare. We used data from seven pre-industrial Finnish populations and  
231 addressed these shortcomings by (1) quantifying variation in early-life disease exposure as  
232 the proportion of children alive in each of the first five years of an individual's life who died  
233 of infectious diseases; (2) assessing links between disease exposure and cause of death; (3)  
234 determining the association between early-life disease exposure and reproduction. Our results  
235 show that individuals experiencing higher early-life disease exposure do not experience  
236 higher all-cause mortality risk or risk of mortality from CVD, stroke or cancer during  
237 adulthood. We also found no support for associations between early-life disease exposure and  
238 reproductive performance in either males or females. Our results do not support the link  
239 between early-life disease exposure and chronic inflammation-associated mortality risk  
240 which forms the crux of the 'cohort morbidity phenotype' hypothesis.

241

242 Our analyses showed that, controlling for individual parish, social class, sex, birth season,  
243 twinning status, birth interval and temporal changes in disease exposure (Fig. 1) there was no  
244 statistical support for an association between early-life disease exposure and mortality risk  
245 after age 15 (Table 1 and SI Appendix, Table S4). Previous studies have found evidence for  
246 links between early-life cohort mortality rate and later survival (4, 5, 8), but without  
247 considering that such a measure of disease exposure could be confounded by spatial  
248 variation, social status, other intrinsic and extrinsic conditions as well as disease exposure,  
249 and improvements in living conditions and hygiene across time. Demographic studies have  
250 applied statistical techniques such as Hodrick-Prescott decomposition to trending

251 environmental data and found generally weak associations between early- and later cohort  
252 mortality (37–40), though some of these studies have still detected significant correlations  
253 (19, 41). By applying Hodrick-Prescott decomposition we were able to account for declines  
254 in death rates from infection across time by using deviations from this trend as our measure  
255 of early-life disease exposure (Fig. 2). Unlike these studies, however, we use a measure of  
256 disease exposure explicitly based on prevalence of deaths from the infectious diseases most  
257 likely to affect infants (SI Appendix, Tables S2 & S3).

258

259 Early-life disease exposure may influence later-life health in a number of ways, and it is this  
260 complexity that may have prevented us from detecting significant associations. For example,  
261 early infections may have negative associations with later health through chronic  
262 physiological ‘scarring’, but may also have positive associations due to acquired immunity or  
263 selection of the most robust individuals who were able to survive the harsh early environment  
264 (42). There may also be changes in the factors which influence mortality in different parts of  
265 the life course, such that the relative importance of scarring and selection may vary with age  
266 (43) and the effects of early adversity may be short-lived, or be negligible in comparison to  
267 events which occur in adult life (39). A study using data from Sweden, collected 1813–1968  
268 and calculating early-life disease risk based on neonatal mortality rates, showed that effects  
269 of exposure on mortality varied with age: individuals experiencing higher neonatal mortality  
270 rates showed higher mortality initially, but lower mortality in later life (44). These results  
271 suggest initial scarring effects of disease at younger ages, followed by a strengthening  
272 influence of acquired immunity and/or selection of robust individuals at later ages. Thus,  
273 rather than being unimportant for later-life survival, early-life infections may have both  
274 negative (through scarring) and positive (through acquired immunity and/or selection)

275 associations with later survival which may combine to culminate in no detectable overall  
276 effect.

277

278 The ‘cohort morbidity phenotype’ (5–8) proposes a link between infection, chronic  
279 inflammation, and risk of CVD, stroke and cancer (14). The exact prevalence of these  
280 diseases in our study population is unknown: although 9.6% of deaths were due to CVD or  
281 stroke, a further 13.7% of deaths were due to ‘old age’ and could have been related to chronic  
282 inflammation (SI Appendix, Table S5). Exposure to infection in early life has also been  
283 linked to respiratory disorders in later life (16, 17, 45), which is noteworthy given that 14.9%  
284 of deaths in our studied individuals were from tuberculosis (SI Appendix, Table S5).  
285 Focusing on deaths from CVD, stroke and cancer, we did not find evidence that early-life  
286 disease exposure was associated with these causes of death, linked to chronic inflammation  
287 by the ‘cohort morbidity phenotype’ (SI Appendix, Table S6). It should be noted that, with  
288 only 346/7283 (5% of) individuals scored as dying of CVD, stroke or cancer, our statistical  
289 power to detect links between early disease exposure and cause of death may have been low.  
290 Meanwhile, studies of contemporary populations experiencing high exposure to infection and  
291 high levels of inflammatory markers rarely show signs of CVD despite their short lifespans,  
292 suggesting that a low level of CVD, stroke and cancer is simply the ancestral state and that  
293 high prevalence of these conditions may be a consequence of longer lifespans in  
294 industrialized populations (46). Finally, it may be possible that adverse effects of chronic  
295 inflammation are only apparent in populations experiencing resource-rich or pathogen-poor  
296 environments, and that in our (relatively) resource-poor and pathogen-rich environments,  
297 heightened inflammatory responses are advantageous or neutral. Our study individuals were  
298 born before 1850, and thus it is possible that our results may not be generalizable to later  
299 cohorts, which experienced increased adult nutrition and healthcare. On the other hand, if

300 chronic inflammation does not lead to CVD, stroke and cancer in low-nutrition, high-  
301 pathogen environments, then the ‘cohort morbidity phenotype’ would not offer an  
302 explanation for longer modern lifespans being due to reduced or delayed instances of these  
303 disorders.

304

305 The majority of studies of early-life disease exposure have focused solely on associations  
306 with mortality, and our results are among the first to address links between early-life disease  
307 exposure and reproductive success. We found, however, that associations between early-life  
308 disease exposure and lifetime reproductive success (LRS), age at first reproduction, lifetime  
309 breeding success and child survival rate were absent in both sexes. Life-history theory  
310 predicts that high mortality risk in early life should lead to earlier age at maturity (21), which  
311 should be associated with higher LRS, predictions which have been supported by population-  
312 (23, 47) and individual-level studies showing that adverse early-life conditions such as  
313 experience of mortality, a deprived upbringing, experience of famine or chronic childhood  
314 illness are associated with earlier age at first birth and higher fertility (22, 23, 25). In contrast,  
315 early-life stressors such as low food availability have been linked to later-life health problems  
316 (48) and, in our study population, reduced reproductive success (49, 50). These effects may  
317 be mediated through neonatal infections which stunt growth and diminish later reproductive  
318 success (26, 27), or maternal infections which reduce the birthweight of focal individuals (28,  
319 29). In a historical German population, birth during a high-mortality period was associated  
320 with lower child survivorship in males and lower lifetime births in females (51). Thus,  
321 adverse conditions in early pre- or post-natal life may slow growth and development, affect  
322 reproductive physiology and postpone first birth (52). As with survival, the conflicting  
323 negative (scarring) and positive (selection) effects of disease exposure may underpin our  
324 finding of no overall association between early-life disease exposure and later-life

325 reproductive success. Alternatively, the lack of association in the present study could indicate  
326 that links detected in previous studies were due to factors other than disease. A recent study  
327 on data from 251 20<sup>th</sup> century English women used path analysis to show that although social  
328 class was linked to age at first birth, and illness in the first year of life was linked to adult  
329 height, there was no direct or indirect effect of childhood illness on age at first birth or LRS  
330 (53). Our results, with N>1000 are therefore unlikely to reflect lack of statistical power and it  
331 may therefore simply be that effects of early-life infections on reproduction are weak and  
332 overshadowed by socio-economic factors.

333

334 Unlike many previous studies testing for links between early-life disease exposure and later-  
335 life fitness, we (1) incorporated data on disease prevalence into our measure of early-life  
336 disease exposure; (2) determined associations between disease exposure and cause of death;  
337 (3) assessed associations between disease exposure and reproductive success; (4) controlled  
338 for several important social factors; (5) de-trended our measure of early-life disease exposure  
339 to account for temporal trends; (6) examined early-life effects not only around birth but up to  
340 age five. An ‘ideal’ study would link full health records from each individual documenting  
341 previous illness with lifetime follow-up on survival and reproduction. Our results should be  
342 interpreted with the acknowledgment that infection of any individual in our population at any  
343 time is unknown. Instead, our measure of disease exposure assumes that a higher child  
344 mortality rate from infection meant that, on average, individuals experiencing such conditions  
345 were more likely to be exposed to infections. In addition, we included the 22% of child  
346 deaths due to ‘unknown diseases’ in the ‘infectious’ category since deaths from infection  
347 were generally more common than those due to other causes, especially in children (Tables  
348 S3 & S5), but this may have introduced noise into our calculation of disease exposure if not  
349 all such unknown diseases were indeed infectious. It therefore remains possible that inability

350 to detect an association between early disease exposure and fitness is due to our way of  
351 calculating early disease exposure. Ability to accurately determine cause of death with poor  
352 medical knowledge is a general caveat of using historical data. These two caveats aside, a  
353 strength is that we have used parish-specific records rather than nationwide figures, thereby  
354 gaining the most geographically-restricted estimate of disease exposure possible.

355

356 The results of our study strongly challenge the hypothesis that links between early and later-  
357 life mortality rates are mediated through effects of early-life infectious diseases on chronic  
358 inflammation. Longitudinal study of contemporary populations will be essential for detecting  
359 the mechanisms underpinning such effects (53); for example, for disentangling the role of  
360 early and later conditions in influencing mortality risk. With the increasing availability of  
361 national, individual-based health data, this endeavour will be achievable in a more explicit  
362 manner than shown here. Our study, however, uses data from 100 cohorts of a pre-industrial  
363 population and controls for variation in social class and birth circumstances to provide a  
364 unique snapshot of how early-life conditions may have influenced human populations for  
365 much of our evolutionary history (6, 7). We conclude, in agreement with several other  
366 authors (37, 39, 40) that the influence of improved adult, rather than childhood, conditions  
367 are likely to have driven recent improvements in life expectancy.

368

## 369 MATERIALS AND METHODS

370

### 371 Study population and data collection

372

373 We investigated associations between early-life disease exposure and later-life survival and  
374 reproductive success using longitudinal data collected from 18<sup>th</sup>-19<sup>th</sup> century church records

375 in seven populations, termed ‘parishes’, located across the southern half of Finland. Prior to  
376 the 1870s, these populations had high birth and death rates, primitive agricultural technology,  
377 unreliable contraception and medical care (31) and were strictly monogamous. Individuals  
378 were divided into two social classes: “Rich” individuals included farm owners, merchants  
379 and craftsmen; “Poor” individuals included crofters and labourers (32).

380

### 381 **Causes of death and early-life disease exposure**

382

383 We have collected data including birth, parentage, marriage, reproduction and death for a  
384 total of 72,564 individuals in the parishes of Hiittinen, Ikaalinen, Jaakkima, Kustavi,  
385 Pulkkila, Rautu and Rymättylä born 1702-2012. We restricted our analyses to 21,539  
386 individuals born after 1750, since before this records were patchy, and born before 1851,  
387 ensuring that we only used data from individuals with natural mortality and fertility (31). We  
388 further restricted our analyses to individuals who survived to at least age 15 (15,237) and  
389 those of known maternal identity and social class (N = 7327). This was done to ensure we  
390 captured longer-term effects of disease exposure; and to control for variation in survival and  
391 reproductive success between families and social classes (32). Finally, because birth order  
392 (34), twinning status (33) and birth season (35) are all linked to fitness in this population, we  
393 restricted our analyses to individuals for whom this information was recorded, leaving a final  
394 sample size of 7283 individuals, born 1751-1850. None of our study individuals moved  
395 between parishes between birth and adulthood.

396

397 Based on previous studies examining links between early exposure to mortality and later-life  
398 health and fitness (19, 38–40), we aimed to estimate an individual’s experience of infectious  
399 diseases in their first five years of life, as illustrated by the prevalence of death from

400 infectious diseases recorded by the church (54). Since children and adults are vulnerable to  
401 different diseases in this population (SI Appendix, Tables S3 & S5), we recorded the number  
402 of individuals aged 0-10 years who were alive in the parish in each year, 1751-1855, a total  
403 of 194,182 person-years (SI Appendix, Table S2), during which time 4780 children died of a  
404 known cause, 73% of which were due to infectious diseases (SI Appendix, Table S3). 22%  
405 of deaths were from an ‘unknown disease’, which we included in the ‘infectious’ category.  
406 We gained a parish-specific measure of ‘disease exposure’ for each year by dividing the  
407 number of children dying of infectious diseases in a given year in a given parish by the  
408 number of children alive in the same parish in the same year (Fig. 1). Thus, our measure of  
409 early-life disease exposure is based on the child mortality rate from infectious diseases during  
410 an individual’s early life. We then de-trended this measure using Hodrick-Prescott  
411 decomposition, and used the deviations from the trend (the ‘cycle’ component) as our  
412 measure of disease exposure (Fig. 2). Early-life disease exposure was multiplied by 100 prior  
413 to use in analysis to aid model convergence.

414

## 415 **Statistical analysis**

416

417 We determined the association between disease exposure in each of the first five years of life  
418 and later-life fitness, controlling for other potentially confounding effects (SI Appendix,  
419 Table S1). All models described below were mixed-effects models with maternal identity and  
420 birth year as random effects, to account for between-family and -cohort variation in fitness.  
421 All models estimated the association between early-life disease exposure and fitness after  
422 controlling for the following fixed effects, which were retained in all models. Parish (seven  
423 levels), social class (Rich or Poor) and sex (Male or Female) were all included as categorical  
424 fixed effects. Since child survival is higher for those born in August, September and October

425 compared to other months (35), birth season was included as a two-level categorical variable  
426 (High or Low survival). First-born males have higher prospects of wealth inheritance,  
427 marriage and reproduction in this population, and individuals born with a longer intervening  
428 period after their previous sibling may have higher fitness (34). We incorporated both of  
429 these effects into a four-level categorical variable, Inter-Birth Interval (IBI: Firstborn; born  
430 <2 years after sibling; born <3 years after sibling; born >3 years after sibling). Finally, all  
431 models accounted for the link between twin status at birth and later-life fitness (33) by  
432 including twin status at birth (singleton or twin) as a categorical fixed effect. All analyses  
433 were conducted in R version 3.2.3.

434

#### 435 **Mortality risk**

436

437 We determined the association between early-life disease exposure and mortality risk after  
438 the age of 15 in 7283 individuals using mixed-effects Cox models in the R package ‘coxme’.  
439 We constructed a ‘base’ model containing the fixed and random effects described above  
440 (model 0) and then built a series of models testing the association between early-life disease  
441 exposure and mortality risk. We considered disease exposure in each of the first five years of  
442 an individual’s life as separate variables: Year 0 (birth year), Year 1, Year 2, Year 3 and Year  
443 4. We first tested for main effects of disease exposure in each of the first five years of life  
444 separately by (models 1-5) comparing the ‘base’ model with separate models containing each  
445 of the first five years’ disease exposure, fitted as covariates, using likelihood ratio test  
446 (LRTs), where the  $\chi^2$ -distributed test statistic is calculated as  $-2*(\text{LogLik}_{\text{model1}} -$   
447  $\text{LogLik}_{\text{model2}})$ . We next (model 6) fitted all five disease exposure years as covariates in the  
448 same model and compared model 6 to model 0 using an LRT. We determined which of  
449 models 1-5 fitted best by comparing their Akaike Information Criterion (AIC) values, where

450 the lowest value provided the best fit (36). We did not encounter any situation where more  
451 than one of models 1-5 outperformed model 0. This allowed us to determine which disease  
452 exposure year was most strongly associated with mortality risk. We took the best-fitting  
453 disease exposure year forwards into model 7, which tested for a non-linear association  
454 between the best-fitting disease exposure year and survival by comparing the model with a  
455 linear term of disease exposure to one with both linear and quadratic terms. Finally, we tested  
456 for interactions involving the best-fitting disease exposure year by comparing the model with  
457 the main effect of the best-fitting disease exposure year to models with interactions between  
458 disease exposure and parish (model 8), social class (model 9) and sex (model 10). We  
459 subsequently restricted the analyses to individuals who lived to at least age 50 (N=3822),  
460 since these individuals are most likely to fall victim to CVD, stroke and cancer.

461

## 462 **Cause of death**

463

464 We next aimed to determine whether early-life disease exposure was associated with  
465 increased risk of death from stroke, CVD or cancer versus other causes. Each individual was  
466 scored as dying from stroke, CVD or cancer (1; 346 individuals) with the remainder (6937  
467 individuals) scored as 0. Cause-of-death analyses were performed using mixed-effects Cox  
468 models, following the same fixed and random effects structure and model selection procedure  
469 as described above. We then repeated this analysis on individuals who survived until at least  
470 the age of 50, since 263/346 (76%) of deaths from these causes occurred in individuals aged  
471 50 and over.

472

## 473 **Lifetime reproductive success**

474

475 We next determined the association between early-life disease exposure and later-life  
476 reproductive success by quantifying lifetime reproductive success (LRS) as the number of  
477 children an individual produced that survived to age 15. From the original sample size of  
478 7283, we removed data from 1200 individuals who were unmarried or whose marriage status  
479 was unknown; 488 individuals who had unknown LRS; 1417 individuals whose reproductive  
480 success was unknown or whose children had unknown fates to age 15; 1333 individuals who  
481 failed to survive to the age of 50 and therefore complete their reproductive lifespan (49). We  
482 therefore focused on fertility differences within marriage to determine differences in  
483 reproductive capacity, rather than ‘mating success’ (or marriage) which is influenced by a  
484 host of social and cultural factors which contribute to finding a partner. The final data set for  
485 analysis of LRS contained 2845 individuals, with 1512 females and 1333 males. The analyses  
486 were separated by sex due to the biological and social differences governing variation in  
487 reproductive success between the sexes in this population. We visually inspected diagnostic  
488 plots of several different modelling structures, including linear mixed-effects models  
489 (LMMs) of untransformed LRS; LMMs of log-transformed LRS; LMMs of Box-Cox  
490 transformed LRS; generalized linear mixed effects models of untransformed LRS with  
491 Poisson, Poisson-lognormal and negative binomial error structures. In both males and  
492 females, Box-Cox transformation provided the most favourable diagnostics, so LRS was  
493 modelled using linear mixed-effects models (LMMs) of  $(LRS+1)^{0.5}$  in the R package ‘lme4’.  
494 The fixed and random effect structures of models 1-9 were the same as for the Cox models  
495 described above, although since we performed separate analysis for males and females, we  
496 did not fit model 10, which included the interaction between early-life disease exposure and  
497 sex.

498

499 Finally, we tested for associations between early-life disease exposure and the reproductive  
500 traits which contribute to LRS. We therefore performed analyses determining whether early-  
501 life disease exposure was associated with later age at first reproduction, reduced lifetime  
502 breeding success (number of children born) and/or child survival rate to age 15. These  
503 analyses are described in full in the SI Appendix.

504

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681 **FIGURE LEGENDS**

682

683 **Figure 1:** Changes in child mortality from infectious disease in each year, 1751-1855 and  
684 calculation of our measure of early-life disease exposure. Child (aged 0-10) mortality rate  
685 from infectious disease varied substantially across the early lives of our study individuals.  
686 Grey points and lines show the proportion of children alive in a given year that died of an  
687 infectious disease; black line shows the ‘trend’ component of Hodrick-Prescott  
688 decomposition of the time series. Data and trends were calculated and are shown separately  
689 for (A) all parishes combined; (B) Hiittinen; (C) Ikaalinen; (D) Jaakkima; (E) Kustavi; (F)  
690 Pulkila; (G) Rautu; (H) Rymattyla.

691

692 **Figure 2:** The deviations from the trend in child mortality from infectious disease in the early  
693 years of our study individuals’ lives, represented as the ‘cycle’ component of the Hodrick-  
694 Prescott decomposition. The parish-specific deviations from the trend, our measure of early-  
695 life disease exposure, varied substantially across cohorts in (A) all parishes combined; (B)  
696 Hiittinen; (C) Ikaalinen; (D) Jaakkima; (E) Kustavi; (F) Pulkila; (G) Rautu; (H) Rymattyla.

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